

Symmetric generalized binomial distributions

H. Bergeron, E. M. F. Curado, J. P. Gazeau, and Ligia M. C. S. Rodrigues

Citation: *J. Math. Phys.* **54**, 123301 (2013); doi: 10.1063/1.4837135

View online: <http://dx.doi.org/10.1063/1.4837135>

View Table of Contents: <http://jmp.aip.org/resource/1/JMAPAQ/v54/i12>

Published by the AIP Publishing LLC.

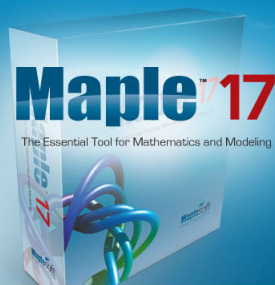
Additional information on J. Math. Phys.

Journal Homepage: <http://jmp.aip.org/>

Journal Information: http://jmp.aip.org/about/about_the_journal

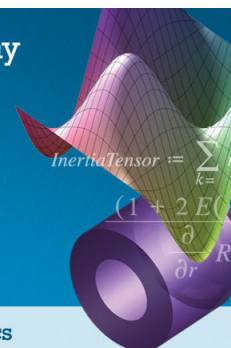
Top downloads: http://jmp.aip.org/features/most_downloaded

Information for Authors: <http://jmp.aip.org/authors>



The most comprehensive support for Physics in any mathematical software package

- State-of-the-art environment for algebraic computations in physics
- Access to Maple's full mathematical power, programming language, visualization routines, and documentation creation tools
- The only system with the ability to handle a wide range of physics computations as well as pencil-and-paper style input and textbook-quality display of results
- A programming library that gives access to almost 100 internal commands to write programs or extend the capabilities of the Physics package



Click to learn more >>

World-leading tools for performing calculations in theoretical physics

Symmetric generalized binomial distributions

H. Bergeron,¹ E. M. F. Curado,^{2,3} J. P. Gazeau,^{2,4}
 and Ligia M. C. S. Rodrigues^{2,a)}

¹Univ Paris-Sud, ISMO, UMR 8214, 91405 Orsay, France

²Centro Brasileiro de Pesquisas Físicas, Rua Xavier Sigaud 150 22290-180 - Rio de Janeiro, Brazil

³Instituto Nacional de Ciência e Tecnologia - Sistemas Complexos, Rua Xavier Sigaud 150, 22290-180 - Rio de Janeiro, RJ, Brazil

⁴APC, UMR 7164, Univ Paris Diderot, Sorbonne Paris Cité, 75205 Paris, France

(Received 30 July 2013; accepted 15 November 2013; published online 9 December 2013)

In two recent articles, we have examined a generalization of the binomial distribution associated with a sequence of positive numbers, involving asymmetric expressions of probabilities that break the symmetry *win-loss*. We present in this article another generalization (always associated with a sequence of positive numbers) that preserves the symmetry *win-loss*. This approach is also based on generating functions and presents constraints of non-negativeness, similar to those encountered in our previous articles. © 2013 AIP Publishing LLC. [<http://dx.doi.org/10.1063/1.4837135>]

I. INTRODUCTION

In two previous papers,^{1,2} we have introduced Bernoulli-like (or binomial-like) distributions associated to arbitrary sequences of positive numbers and we have given a set of interesting properties. These distributions break the *win-loss* symmetry of the ordinary binomial distribution. In general, they are formal because they are not always positive. We have proved in Ref. 1 that for a certain class of such sequences the corresponding Bernoulli-like distributions are nonnegative and possess a Poisson-like limit law. We also discussed the question of non-negativeness in the general case and we examined under which conditions a probabilistic interpretation can be given to the proposed binomial-like distribution. In Ref. 2, we gave a different approach based on generating functions that allowed to avoid the previous difficulties: the constraints of non-negativeness were automatically fulfilled and a great number of analytical examples became available.

Our previous studies were mainly motivated by the construction and applications of nonlinear coherent states (Ref. 3, see also Refs. 4 and 5 and references therein). In the present work, we give a different generalization of the binomial distribution that preserves the symmetry *win-loss* and which is more appropriate for the analysis of strongly correlated systems and their entropy behavior, as is done, for instance, in Ref. 6. In that paper, the authors present a generalized stochastic model yielding q -Gaussians for all ($q \neq 1$). (The q -Gaussian is a probability distribution, arising from the maximization of the Tsallis entropy under appropriate constraints, which is suitable for strongly correlated random variables⁷). This is achieved by using the Laplace-de Finetti representation theorem for the Leibniz triangle rule,⁸ which embodies strict scale-invariance of interchangeable random variables. They demonstrate that strict scale invariance mandates the associated extensive entropy to be Boltzmann-Gibbs (BG). By extensive we mean that the entropy is proportional to the size, defined in a certain sense, of the system. Our symmetric distributions do not in general fulfill the Leibniz triangle rule. Indeed, we prove that the only symmetric distributions, besides the ordinary binomial case, which obey this rule are precisely those derived from a q -exponential. On the other side, we observe that in many other cases this rule is obeyed asymptotically as n goes to infinity. Finally, in the case of the q -exponential we show that the corresponding BG entropy is

^{a)}Electronic addresses: herve.bergeron@u-psud.fr, evaldo@cbpf.br, gazeau@apc.univ-paris7.fr, and ligia@cbpf.br

extensive; in the cases the Leibniz rule is asymptotically verified, the BG entropy is asymptotically extensive.

The questions of non-negativeness encountered in our previous works are also present in the symmetric case and they are approached through a similar procedure. Interestingly, the characterization of generating functions compatible with the constraints of non-negativeness leads to the set of solutions already found in Ref. 2.

A potential application of our results concerns the notion of bipartite entanglement characterized by approaches involving the Tsallis q -entropy.⁹ Another application concerns the construction of coherent states built from these symmetric deformed binomial distribution, in the same way spin coherent states are built from ordinary binomial distribution.⁵

In Sec. II, we review the asymmetric generalized binomial distributions, the associated generating functions, and some useful definitions from Refs. 1 and 2. We develop in Sec. III the case of symmetric generalized binomial distributions with the necessary mathematical tools. They lead to a characterization of the solutions compatible with a complete probabilistic interpretation. The generalized Poisson distribution is introduced as well. In Sec. IV, we derive the general expression for expectation values and moments of generalized binomial distributions in the symmetric case. In Sec. V, we present some analytical and numerical examples that illustrate the theoretical framework developed in Sec. III. These examples belong to a class that we define as *logarithmic scale invariant generating functions*.

In Sec. VI, we present “nonlinear spin” coherent states built from those symmetric deformed binomial distributions, and suggest possible physical applications of these deformed distributions. In Sec. VII, we present our final comments. Technical details and relevant results with proofs are given in Appendixes A–E.

II. THE ASYMMETRIC GENERALIZED BINOMIAL DISTRIBUTION: A BRIEF REVIEW

For a sequence of n independent trials with two possible outcomes, “win” and “loss,” the probability of obtaining k wins is given by the binomial distribution $p_k^{(n)}(\eta)$

$$p_k^{(n)}(\eta) = \binom{n}{k} \eta^k (1 - \eta)^{n-k} = \frac{n!}{(n-k)!k!} \eta^k (1 - \eta)^{n-k}, \quad (1)$$

where the parameter $0 \leq \eta \leq 1$ is the probability of having the outcome “win” and $1 - \eta$, the outcome “loss.” Therefore, we can say that the Bernoulli binomial distribution above corresponds to the sequence of non-negative integers $n \in \mathbb{N}$.

In two recent articles,^{1,2} we generalized the binomial law¹⁰ using an increasing sequence of nonnegative real numbers $\{x_n\}_{n \in \mathbb{N}}$, where by convention $x_0 = 0$. The new “formal” (i.e., not necessarily positive) probabilities $\mathfrak{p}_k^{(n)}(\eta)$ are defined as

$$\mathfrak{p}_k^{(n)}(\eta) = \frac{x_n!}{x_{n-k}!x_k!} \eta^k p_{n-k}(\eta) \equiv \binom{x_n}{x_k} \eta^k p_{n-k}(\eta), \quad (2)$$

where the factorials $x_n!$ are given by

$$x_n! = x_1 x_2 \dots x_n, \quad x_0! \stackrel{\text{def}}{=} 1, \quad (3)$$

and where the polynomials $p_n(\eta)$ are constrained by the normalization

$$\forall n \in \mathbb{N}, \quad \sum_{k=0}^n \mathfrak{p}_k^{(n)}(\eta) = 1. \quad (4)$$

Note that this distribution is called asymmetric in the sense that the expression (2) is not invariant under $k \rightarrow n - k$ and $\eta \rightarrow 1 - \eta$, as it happens in the binomial case. The polynomials $p_n(\eta)$ can be obtained by recurrence from Eq. (4), the first ones being $p_0(\eta) = 1$ and $p_1(\eta) = 1 - \eta$. The $p_n = \mathfrak{p}_0^{(n)}$ possess a probabilistic interpretation as long as $\mathfrak{p}_k^{(n)}$ are themselves probabilities. This latter condition holds only if the $p_n(\eta)$ are nonnegative. Therefore, non-negativeness of the p_n for $\eta \in [0, 1]$ is a key-point to preserve the statistical interpretation. In this case, the quantity $\mathfrak{p}_k^{(n)}(\eta)$ can

still be interpreted as the probability of having k wins and $n - k$ losses in a sequence of *correlated* n trials. But, due to the asymmetry, there can be bias favoring either win or loss, even in the case $\eta = 1/2$.

In order to analyze this positiveness constraint and characterize the solutions, an analysis based on generating functions has been developed in Ref. 2.

In Ref. 1, it was proved that the analytical function

$$\mathcal{N}(t) = \sum_{n=0}^{\infty} \frac{t^n}{x_n!} \quad (5)$$

gives the generating function of $p_k(\eta)$ (independently of the non-negativeness question) thanks to

$$\frac{\mathcal{N}(t)}{\mathcal{N}(\eta t)} = \sum_{k=0}^{\infty} p_k(\eta) \frac{t^k}{x_k!} \quad (6)$$

we defined in Ref. 2 the set Σ as the set of entire series $\mathcal{N}(t) = \sum_{n=0}^{\infty} a_n t^n$ possessing a non-vanishing radius of convergence and verifying $a_0 = 1$ and $\forall n \geq 1, a_n > 0$. We associate to each of these functions \mathcal{N} the positive sequence $\{x_n\}_{n \in \mathbb{N}}$ defined as $x_0 = 0$ and $\forall n \geq 1, x_n = a_{n-1}/a_n = n \mathcal{N}^{(n-1)}(0)/\mathcal{N}^{(n)}(0)$. Then we obtain $a_n = 1/(x_n!)$ and the definition of Eq. (5) is recovered. We also defined Σ_0 as the set of entire series $f(z) = \sum_{n=0}^{\infty} a_n z^n$ possessing a non-vanishing radius of convergence and verifying the conditions $a_0 = 0, a_1 > 0$ and $\forall n \geq 2, a_n \geq 0$. Finally, we proved that

$$\Sigma_+ := \{\mathcal{N} \in \Sigma \mid \forall \eta \in [0, 1], p_n(\eta) > 0\} = \{e^F \mid F \in \Sigma_0\}; \quad (7)$$

Σ_+ is the set of generating functions solving the non-negativeness problem.

An interesting property reinforcing the probabilistic interpretation is the Poisson-like limit of the asymmetric deformed binomial distribution. If the sequence $\{x_n\}_{n \in \mathbb{N}}$ is such that, at fixed finite m , $\lim_{n \rightarrow \infty} \frac{x_n - m}{x_n} = 1$, then the limit when $n \rightarrow \infty$ of $p_k^{(n)}(\eta)$ with $\eta = t/x_n$ is equal to $\frac{1}{\mathcal{N}(t)} \frac{t^k}{x_k!}$.

III. THE SYMMETRIC GENERALIZED BINOMIAL DISTRIBUTION

Using the notations of Sec. II, we propose to study the following generalization of the binomial distribution:

$$p_k^{(n)}(\eta) = \frac{x_n!}{x_{n-k}! x_k!} q_k(\eta) q_{n-k}(1 - \eta), \quad (8)$$

where the $\{x_n\}$ is a non-negative sequence and the $q_k(\eta)$ are polynomials of degree k , while η is a running parameter on the interval $[0, 1]$. The $p_k^{(n)}(\eta)$ are constrained by:

(a) the normalization

$$\forall n \in \mathbb{N}, \quad \forall \eta \in [0, 1], \quad \sum_{k=0}^n p_k^{(n)}(\eta) = 1, \quad (9)$$

(b) the non-negativeness condition (due to statistical interpretation)

$$\forall n, k \in \mathbb{N}, \quad \forall \eta \in [0, 1], \quad p_k^{(n)}(\eta) \geq 0. \quad (10)$$

The normalization condition (9) applied to the case $n = 0$ leads to

$$\forall \eta \in [0, 1], \quad p_0^{(0)}(\eta) = q_0(\eta) q_0(1 - \eta) = 1. \quad (11)$$

Since q_0 is assumed to be a polynomial of degree 0, we deduce that $q_0(\eta) = \pm 1$. *In all the remainder, we assume*

$$q_0(\eta) = 1. \quad (12)$$

Using this value of q_0 we deduce

$$\forall n \in \mathbb{N}, \quad \forall \eta \in [0, 1], \quad p_0^{(n)}(\eta) = q_n(1 - \eta). \quad (13)$$

Therefore, the non-negativeness condition (10) is equivalent to the non-negativeness of the polynomials q_n on the interval $[0, 1]$. Here, like in the asymmetric case, the quantity $p_k^{(n)}(\eta)$ can be interpreted as the probability of having k wins and $n - k$ losses in a sequence of *correlated* n trials. Besides, as we recover the invariance under $k \rightarrow n - k$ and $\eta \rightarrow 1 - \eta$ of the binomial distribution, no bias (in the case $\eta = 1/2$) can exist favoring either win or loss.

A. Generating functions

Setting aside the non-negativeness condition, different generating functions (and different sets of polynomials) can be associated with the conditions expressed above. Indeed, given a sequence $\{x_n\}$ and its generalized exponential $\mathcal{N}(t)$ introduced in (5), we define the related generating function as

$$F(\eta; t) := \sum_{n=0}^{\infty} \frac{q_n(\eta)}{x_n!} t^n. \quad (14)$$

Due to the normalization property (9), it is easy to show that

$$F(\eta; t) F(1 - \eta; t) = \mathcal{N}(t), \quad (15)$$

which is a nonlinear functional equation for the variable η with a given boundary condition. The general solution is found from Ref. 11

$$F(\eta; t) = \pm \sqrt{\mathcal{N}(t)} e^{\Phi(\eta, 1 - \eta; t)}, \quad (16)$$

where the arbitrary function $\Phi(x, y; t)$ is antisymmetric with respect to the permutation of variables x and y , $\Phi(x, y; t) = -\Phi(y, x; t)$.

B. The simplest class of generating functions

The simplest generating function giving rise to polynomials of binomial type¹² corresponds to the choice $\Phi(x, y; t) = (x - y)\varphi(t)$, together with positive sign in (16) and the boundary condition

$$F(0; t) = 1 \Leftrightarrow F(1; t) = \mathcal{N}(t). \quad (17)$$

This leads to $\mathcal{N}(t) = e^{2\Phi(1, 0; t)} = e^{2\varphi(t)}$ and eventually to

$$F(\eta; t) = e^{(2\eta)\varphi(t)} = (\mathcal{N}(t))^\eta. \quad (18)$$

Hence, starting from Eq. (5) and a x_n -generating function $\mathcal{N}(t) \in \Sigma$ (the definition of Σ was given above), let us define the functions q_n as

$$\forall \eta \in [0, 1], \quad \mathcal{N}(t)^\eta = \sum_{n=0}^{\infty} q_n(\eta) \frac{t^n}{x_n!}. \quad (19)$$

This definition makes sense since:

- (a) $\mathcal{N} \in \Sigma$ implies that \mathcal{N} is analytical around $t = 0$,
- (b) $\mathcal{N} \in \Sigma$ implies $\mathcal{N}(0) = 1$.

Then $\ln(\mathcal{N}(t))$ is analytical around $t = 0$, and therefore $\mathcal{N}(t)^\eta = \exp(\eta \ln(\mathcal{N}(t)))$ possesses a convergent series expansion around $t = 0$ (for all $\eta \in \mathbb{C}$). The functions q_n defined by (19) verify the following properties, whose proof is given in Appendix A:

- (a) The two first polynomials $q_n(\eta)$ are $q_0(\eta) = 1$ (according to our choice given by Eq. (12)) and

$$q_1(\eta) = \eta; \quad (20)$$

more generally the q_n verify the recurrence relation

$$\forall n \in \mathbb{N}, \forall \eta \in [0, 1], \quad q_{n+1}(\eta) = \eta \frac{x_{n+1}}{n+1} \sum_{k=0}^n \binom{x_n}{x_k} \frac{n-k+1}{x_{n-k+1}} q_k(\eta-1); \quad (21)$$

(b) the q_n 's are polynomials of degree n obeying

$$\forall n \in \mathbb{N}, q_n(1) = 1, q_n(0) = 0, \forall n \neq 0, \tag{22}$$

and they fulfill the normalization condition (9);

(c) the q_n 's fulfill the functional relation

$$\forall z_1, z_2 \in \mathbb{C}, \forall n \in \mathbb{N}, \sum_{k=0}^n \binom{x_n}{x_k} q_k(z_1) q_{n-k}(z_2) = q_n(z_1 + z_2). \tag{23}$$

We give in Appendix B a combinatorial expression of the polynomials q_n issued from the obvious formula

$$q_n(\eta) = \frac{x_n!}{n!} \frac{d^n}{dt^n} (\mathcal{N}(t))^\eta \Big|_{t=0}. \tag{24}$$

We note that these polynomials, suitably normalized, are of binomial type.¹² Indeed, with the definition $\tilde{q}_n(\eta) = \frac{n!}{x_n!} q_n(\eta)$, we have $\tilde{q}_n(z_1 + z_2) = \sum_{k=0}^n \binom{n}{k} \tilde{q}_k(z_1) \tilde{q}_{n-k}(z_2)$.

We deduce from (b) that the series expansion in t of the function $\mathcal{N}(t)^\eta$ generates a family of polynomials q_n that fulfill one half of the sought conditions: it remains to specify under which conditions the non-negativeness problem can be solved. Note that the usual binomial distribution is included in this family: it corresponds to the obvious choice $\mathcal{N}(t) = e^t$.

C. The non-negativeness problem

The procedure is similar to the one developed in Ref. 2.

Since we already know that $q_0(\eta) = 1$ and $\forall n \neq 0, q_n(0) = 0$, the non-negativeness condition is equivalent to specify that for any $\eta \in]0, 1]$, $q_n(\eta) > 0$ and then the function $t \mapsto \mathcal{N}(t)^\eta$ belongs to Σ . We prove in Appendix C that the set $\tilde{\Sigma}_+$, defined as

$$\tilde{\Sigma}_+ := \{ \mathcal{N} \in \Sigma \mid \forall \eta \in]0, 1], q_n(\eta) > 0 \}, \tag{25}$$

is in fact

$$\tilde{\Sigma}_+ = \{ e^F \mid F \in \Sigma_0 \} = \Sigma_+. \tag{26}$$

$\tilde{\Sigma}_+$ is the set of generating functions solving the non-negativeness problem for the symmetric case; it is actually the same Σ_+ as for the asymmetric case.

Equation (26) characterizes those generating functions in the simplest class which yield symmetric deformations of the binomial distribution preserving a statistical interpretation. Therefore, all coefficients $q_n^{(k)}(0)$ of the polynomials $q_n(\eta)$ are in fact positive, and then the q_n are positive on \mathbb{R}^+ .

Furthermore, if $\mathcal{N} = e^F$ with $F(t) = \sum_{n=1}^\infty a_n t^n \in \Sigma_0$, the $x_n!$ and the q_n can be generally expressed in terms of the complete Bell polynomials $B_n(a_1, a_2, \dots, a_n)$ ¹³ as

$$x_n! = \frac{n!}{B_n(a_1, a_2, \dots, a_n)} \quad \text{and} \quad q_n(\eta) = \frac{B_n(\eta a_1, \eta a_2, \dots, \eta a_n)}{B_n(a_1, a_2, \dots, a_n)}, \tag{27}$$

thanks to the relation

$$e^{F(t)} = \sum_{n=0}^\infty B_n(a_1, a_2, \dots, a_n) \frac{t^n}{n!}. \tag{28}$$

Moreover, using the Bell polynomials properties¹³ we also have from Eq. (27)

$$q_0(\eta) = 1 \quad \text{and} \quad \forall n \geq 1, q_n(\eta) = \sum_{k=1}^n \frac{B_{n,k}(a_1, a_2, \dots, a_{n-k+1})}{B_n(a_1, a_2, \dots, a_n)} \eta^k, \tag{29}$$

where the $B_{n,k}(a_1, a_2, \dots, a_{n-k+1})$ are the so-called partial Bell polynomials.

Starting from any sequence of positive numbers a_1, a_2, \dots such that $a_1 \neq 0$, Eqs. (27)–(29) give an extended set of explicit solutions of our problem, letting aside all convergence conditions for entire series.

IV. EXPECTED VALUES AND MOMENTS

A. Expectation value

Using successively

$$\mathcal{N}(zt)^\eta \mathcal{N}(t)^{1-\eta} = \sum_{n=0}^{\infty} \frac{t^n}{x_n!} \left(\sum_{k=0}^n z^k \mathfrak{p}_k^{(n)}(\eta) \right), \quad (30)$$

$$(z\partial_z)^K [\mathcal{N}(zt)^\eta \mathcal{N}(t)^{1-\eta}]_{z=1} = \sum_{n=0}^{\infty} \frac{t^n}{x_n!} \langle k^K \rangle_n, \quad (31)$$

we find the simple expression for the expectation value of the k variable

$$\langle k \rangle_n(\eta) = \eta n. \quad (32)$$

We note that we obtain the same value as for the ordinary binomial case.

B. Variance

From (C4) and (31), we derive the formula for the second moment

$$\langle k^2 \rangle_n(\eta) = \eta n^2 + \eta(\eta - 1)c_n, \quad (33)$$

with

$$c_n = \sum_{m=1}^{n-1} m(n-m) \binom{x_n}{x_m} q'_m(0), \quad n \geq 2 \quad c_0 = 0 = c_1, \quad (34)$$

and

$$t^2 \frac{\mathcal{N}'(t)^2}{\mathcal{N}(t)} = \sum_{n=0}^{\infty} \frac{c_n}{x_n!} t^n. \quad (35)$$

Assuming $\mathcal{N}(t) = \exp(\sum_{n=1}^{\infty} a_n t^n)$ and using (27), we also obtain, in terms of Bell polynomials, that

$$c_n = \frac{n!}{B_n(a_1, a_2, \dots, a_n)} \sum_{m=1}^{n-1} \frac{m}{(n-m-1)!} a_m B_{n-m}(a_1, a_2, \dots, a_{n-m}). \quad (36)$$

Thus, the variance is given by

$$(\sigma_k)_n^2(\eta) = \eta(1-\eta)(n^2 - c_n), \quad (37)$$

which entails the interesting property $c_n \leq n^2$ for all n .

V. EXAMPLES WITH LOGARITHMIC SCALE INVARIANCE

A. Logarithmic scale invariance

Suppose that the sequence $\{x_n\}$ depends on a parameter α , $x_n \equiv x_n(\alpha)$, in such a way that for the primary generating function

$$\mathcal{N}(t) \equiv \mathcal{N}(t, \alpha),$$

we have

$$(\mathcal{N}(t, \alpha))^\eta = \mathcal{N}(\eta t, \eta \alpha), \quad (38)$$

i.e., its logarithm is homogeneous of degree 1, $\ln(\mathcal{N}(\eta t, \eta \alpha)) = \eta \ln(\mathcal{N}(t, \alpha))$. Choosing $\eta = 1/\alpha$ allows us to write also

$$\mathcal{N}(t, \alpha) = \mathcal{N}(t/\alpha, 1). \quad (39)$$

The logarithmic scale invariance allows to give by comparison the interesting expression for polynomial q_n

$$q_n(\eta, \alpha) = \eta^n \frac{x_n(\alpha)!}{x_n(\eta\alpha)!}. \quad (40)$$

For the sake of more simplicity in the notation, from now on we will write $q_n(\eta)$ instead of $q_n(\eta, \alpha)$.

This property is shared by the 3 examples studied below. It is interesting to remark that the entropy, which is also a homogeneous function of degree one of its natural variables, can be written in the microcanonical ensemble as $S = \ln \Omega$, where Ω is the number of microscopic states of the system. This suggests that one could interpret $\mathcal{N}(t)$ as the number of states of the system, a number which is modified by the presence of correlations among the events.

B. First example: “ q -exponential”

Let us consider the following family of functions belonging to $\tilde{\Sigma}_+$:

$$\mathcal{N}(t) = \left(1 - \frac{t}{\alpha}\right)^{-\alpha}, \quad \alpha > 0, \quad 0 < t < \alpha \quad (41)$$

that are q -exponentials, by definition $e_q(x) = [1 + (1 - q)x]^{1/(1-q)}$, where the parameter $q = 1 + 1/\alpha$.⁷ We first note that if $\alpha \rightarrow \infty$ then $\mathcal{N}(t) \rightarrow e^t$, i.e., we return to the ordinary binomial case. The corresponding sequence is bounded by α and given by

$$x_n = \frac{n\alpha}{n + \alpha - 1}, \quad \lim_{n \rightarrow \infty} x_n = \alpha. \quad (42)$$

For the factorial, we have

$$x_n! = \alpha^n \frac{\Gamma(\alpha)n!}{\Gamma(n + \alpha)} = \frac{\alpha^n n!}{(\alpha)_n} = \frac{\alpha^n}{\binom{n + \alpha - 1}{n}}, \quad (43)$$

where $(z)_n = \Gamma(z + n)/\Gamma(z)$ is the Pochhammer symbol. The corresponding polynomials are given by

$$q_n(\eta) = \frac{\Gamma(\alpha)}{\Gamma(n + \alpha)} \frac{\Gamma(n + \alpha\eta)}{\Gamma(\alpha\eta)} = \frac{(\alpha\eta)_n}{(\alpha)_n}, \quad (44)$$

and satisfy the recurrence relation

$$q_n(\eta) = \frac{n + \alpha\eta - 1}{n + \alpha - 1} q_{n-1}(\eta), \quad \text{with } q_0(\eta) = 1. \quad (45)$$

In particular $q_1(\eta) = \eta$. The distribution $\mathfrak{p}_k^{(n)}(\eta)$ defined by these polynomials is given by

$$\begin{aligned} \mathfrak{p}_k^{(n)}(\eta) &= \binom{n}{k} \frac{\Gamma(\alpha)}{\Gamma(\alpha + n)} \frac{\Gamma(\eta\alpha + k)}{\Gamma(\eta\alpha)} \frac{\Gamma((1 - \eta)\alpha + n - k)}{\Gamma((1 - \eta)\alpha)} \\ &= \frac{\binom{\eta\alpha + k - 1}{k} \binom{(1 - \eta)\alpha + n - k - 1}{n - k}}{\binom{\alpha + n - 1}{n}}. \end{aligned} \quad (46)$$

It is precisely the Pólya distribution,^{14,15} called also “Markov-Pólya” or “inverse hypergeometric” and more. It was considered by Pólya (1923) in the following urn scheme.¹⁶ From a set of b black balls and r red balls contained in an urn, one extracts one ball and return it to the urn together with c balls of the same color. The probability to have in the urn k black balls after the n th trial is given by (47) with

$$\eta = \frac{b}{b + r}, \quad \alpha = \frac{b + r}{c}, \quad (48)$$

which holds for rational parameters η and α .

Returning to the general distribution (46), the expression of c_n is found to be

$$c_n = n(n-1) \frac{\alpha}{1+\alpha}. \quad (49)$$

Therefore, besides the mean value $\langle k \rangle_n = \eta n$, we have the simple expression for the variance

$$(\sigma_k)_n^2(\eta) = n^2 \eta (1-\eta) \frac{1+\alpha/n}{1+\alpha}. \quad (50)$$

As expected, we check that at large α the behavior of σ_k is ordinary binomial. We also note that it becomes proportional to the mean value at large values of n .

1. Leibniz triangle rule

We prove in Appendix D that the q -exponential is the unique case for which the Leibniz triangle rule holds exactly true. Precisely, the relation

$$\varpi_k^{n-1} = \varpi_k^n + \varpi_{k+1}^n \quad (51)$$

is verified by the quantities, and only by them,

$$\varpi_k^n(\eta) := \frac{p_k^{(n)}(\eta)}{\binom{n}{k}} = \binom{y_n}{y_k} q_k(\eta) q_{n-k}(1-\eta) = \frac{(\eta\alpha)_k ((1-\eta)\alpha)_{n-k}}{(\alpha)_n}, \quad (52)$$

with $y_n := x_n/n$. These relations generalize the r_k^n 's of Ref. 6. The latter correspond to the limit case $\alpha \rightarrow \infty$.

2. Entropy

Following Ref. 6, the above result (51) implies that the entropy which is extensive when $n \rightarrow \infty$ is, like in the ordinary binomial case, Boltzmann-Gibbs. By extensivity we mean that the behavior of the entropy of n events (that are correlated in this example), as a function of n when $n \rightarrow \infty$, is linear. The fact that the entropy which is linear with n is the Boltzmann-Gibbs entropy means that the correlations among the events in this example are not strong enough to change the form of the entropy which is extensive for the case of non-correlated events, i.e., the Boltzmann-Gibbs entropy. Let us explore in a more comprehensive way this statistical aspect of the considered distribution. A natural definition of the BG entropy for the distribution (46), using (52), is

$$S_{\text{BG}}^{(n)} = - \sum_{k=0}^n \binom{n}{k} \varpi_k^n(\eta) \ln \varpi_k^n(\eta). \quad (53)$$

The presence of the binomial coefficient in the sum (53) means a counting of the degeneracy. Its extensive property is shown in Figure 1 where it is numerically compared with the Tsallis entropy $S_q^{(n)} \stackrel{\text{def}}{=} (1 - \sum_{k=0}^n \binom{n}{k} (\varpi_k^n)^q) / (q-1)$.

3. Limit distribution

It is interesting to check (numerically) the limit distribution of (46), normalized by its maximum, in function of k/n , for the example of Fig. 1, i.e., for the values $\alpha = 3$ and $\eta = 1/2$, as $n \rightarrow \infty$, as is shown in Figure 2. There are three curves in this graph, one for $n = 1000$, other for $n = 2000$, and the third one for the analytical expression $2\sqrt{x(1-x)}$. These numerical results strongly suggest that the limit normalized distribution is given by the expression $2\sqrt{x(1-x)}$, which is the Wigner distribution having support $\in [0, 1]$ and centered at $1/2$ (this expression can be obtained from the Wigner distribution $\rho(y) = (8/\pi)\sqrt{(1/2)^2 - y^2}$ changing $y \rightarrow x - 1/2$ and normalizing by its maximum, $2/\pi$). An analytical derivation of this limit distribution deserves interest and we are working on that.

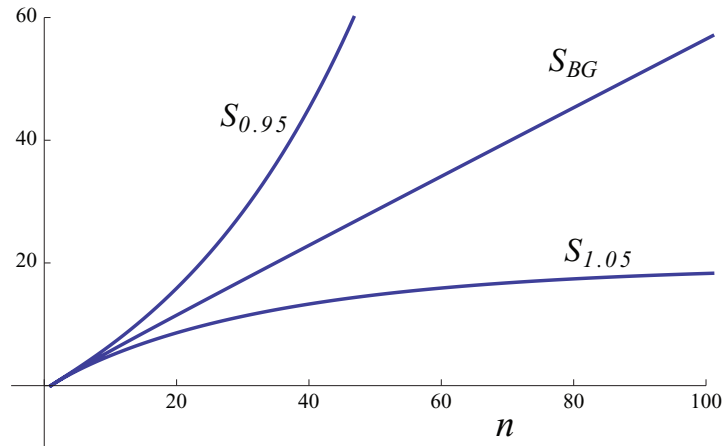


FIG. 1. n -dependence of entropies for the “ q -exponential” case ($\mathcal{N}(t) = (1 - t/\alpha)^{-\alpha}$). The Boltzmann-Gibbs (BG) and Tsallis (S_q) entropies of the distribution are shown for $\eta = 1/2$ and $\alpha = 3$. Upper curve is for $q = 0.95$. Bottom curve is for $q = 1.05$.

C. Second example: Modified Abel polynomials

$$\mathcal{N}(t) = e^{-\alpha W(-t/\alpha)}, \quad \alpha > 0, \tag{54}$$

where W is the Lambert function, i.e., solving the functional equation $W(t)e^{W(t)} = t$. We first note that if $\alpha \rightarrow \infty$ then $\mathcal{N}(t) \rightarrow e^t$. The corresponding sequence is bounded by α/e and given by

$$x_n = \frac{n\alpha}{n + \alpha} \left(1 - \frac{1}{n + \alpha}\right)^{n-2}, \quad \lim_{n \rightarrow \infty} x_n = \alpha/e. \tag{55}$$

We also note that $x_n \rightarrow n$ as $\alpha \rightarrow \infty$. The corresponding factorial is

$$x_n! = n! \frac{\alpha^{n-1}}{(n + \alpha)^{n-1}}. \tag{56}$$

The polynomials q_n 's are given by

$$q_n(\eta) = \eta \frac{\left(\eta + \frac{n}{\alpha}\right)^{n-1}}{\left(1 + \frac{n}{\alpha}\right)^{n-1}}. \tag{57}$$

We verify that $q_0(\eta) = 1$ and $q_1(\eta) = \eta$.

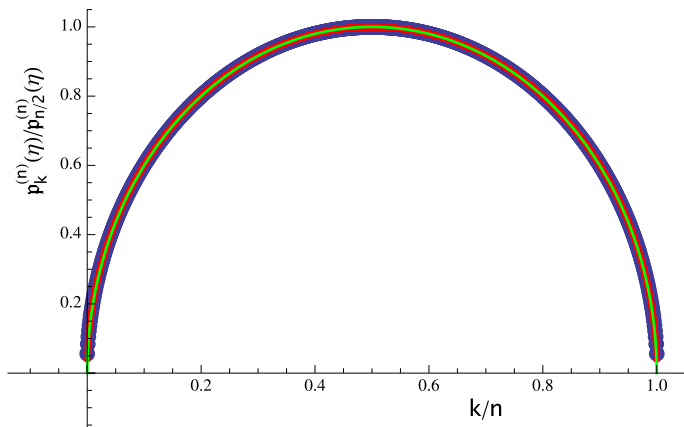


FIG. 2. Limit normalized distribution of the distribution (46) in function of k/n for $\alpha = 3$ and $\eta = 1/2$.

The corresponding probability distribution is given by

$$p_k^{(n)}(\eta) = \binom{n}{k} \eta(1-\eta) \frac{(\eta + k/\alpha)^{k-1} (1-\eta + (n-k)/\alpha)^{n-k-1}}{(1+n/\alpha)^{n-1}}. \tag{58}$$

In order to calculate the variance, we need the explicit form of c_n defined in (34). We first note that

$$q_n'(0) = \frac{n^{n-1}}{(n+\alpha)^{n-1}}.$$

There results for c_n

$$\begin{aligned} c_n &= n! \frac{\alpha}{(n+\alpha)^{n-1}} \sum_{m=0}^{n-2} \frac{(m+1)^m}{m!} \frac{(n-1+\alpha-m)^{n-2-m}}{(n-2-m)!} \\ &= n(n-1) \frac{\alpha}{(n+\alpha)^{n-1}} \mathcal{T}_{n-2}(\alpha+2), \end{aligned} \tag{59}$$

where the polynomials $\mathcal{T}_n(z)$ are defined by the generating function

$$\frac{1}{(1+W(x))^2} e^{-zW(x)} = \sum_{n=0}^{\infty} \frac{\mathcal{T}_n(z)}{n!} (-x)^n. \tag{60}$$

Interesting expressions and properties of the $\mathcal{T}_n(z)$ are given in Ref. 19, particularly related to combinatorics (e.g., subfactorials or derangement numbers). In particular, we have

$$\mathcal{T}_n(z) = \sum_{k=0}^n \frac{n!}{k!} (z+n)^k, \tag{61}$$

and we prove its relation to the incomplete Gamma function¹⁸

$$\mathcal{T}_n(z) = e^{z+n} \Gamma(n+1, z+n), \tag{62}$$

where

$$\Gamma(a, x) = \int_x^{\infty} t^{a-1} e^{-t} dt, \quad \text{Re}(a) > 0. \tag{63}$$

The expression for the variance is then given by

$$\begin{aligned} (\sigma_k)_n^2(\eta) &= n^2 \eta(1-\eta) \left(1 - \frac{n-1}{n} \frac{\alpha}{(n+\alpha)^{n-1}} \mathcal{T}_{n-2}(\alpha+2) \right) \\ &= n^2 \eta(1-\eta) \left(1 - \frac{n-1}{n} \frac{\alpha}{(n+\alpha)^{n-1}} \sum_{k=0}^{n-2} \frac{(n-2)!}{k!} (n+\alpha-2)^k \right). \end{aligned} \tag{64}$$

From (61), we check that at large α the behavior of σ_k is ordinary binomial. We also note that it becomes proportional to the mean value at large values of n .

In this case, we note that the Leibniz triangle is satisfied asymptotically as $n \rightarrow \infty$.

D. Third example: Hermite polynomials

$$\mathcal{N}(t) = e^{t+\frac{a}{2}t^2}, \quad 0 < a < 1. \tag{65}$$

The logarithmic scale invariance holds with $\alpha := 1/a$. The corresponding sequence x_n has the same factorial as in the asymmetric case²

$$x_n! = \left[\frac{i^n \left(\frac{a}{2}\right)^{n/2}}{n!} H_n \left(\frac{-i}{\sqrt{2a}} \right) \right]^{-1} = \left[\sum_{m=0}^{\lfloor \frac{n}{2} \rfloor} \frac{(a/2)^m}{m!(n-2m)!} \right]^{-1} := \frac{1}{\varphi_n(a)}. \tag{66}$$

In particular, $x_1! = x_1 = 1$, $x_2! = 2/(a + 1)$. Also, $x_n = \varphi_{n-1}(a)/\varphi_n(a)$, and we know from Ref. 2 that $x_n \approx \sqrt{n/a}$ as $n \rightarrow \infty$. The corresponding polynomials and probability distributions are, respectively, given by

$$q_n(\eta) = \frac{x_n!}{n!} \left(i \sqrt{\frac{a\eta}{2}} \right)^n H_n \left(-i \sqrt{\frac{\eta}{2a}} \right) \tag{67}$$

and

$$p_k^{(n)}(\eta) = \eta^k (1 - \eta)^{n-k} \frac{\varphi_k(a/\eta) \varphi_{n-k}(a/(1 - \eta))}{\varphi_n(a)}. \tag{68}$$

We know from the general case that $q'_0 = 0$ and $q'_1(0) = 1$. The expression of $q'_2(0)$ is found to be

$$q'_2(0) = \frac{a}{1 + a},$$

and we have $q'_n(0) = 0$ for all $n \geq 3$. Therefore, we have for the c_n 's, $n \geq 2$, the following expressions:

$$c_2 = \frac{2}{1 + a}, \tag{69}$$

$$\begin{aligned} c_n &= n(n - 1) \left[1 + 2 \frac{H_{n-2} \left(\frac{-i}{\sqrt{2a}} \right)}{H_n \left(\frac{-i}{\sqrt{2a}} \right)} \right] \\ &= n(n - 1) \left[1 - \frac{a}{n(n - 1)} \frac{\varphi_{n-2}(a)}{\varphi_n(a)} \right], \quad n \geq 3. \end{aligned} \tag{70}$$

The expression for the variance is then given by

$$(\sigma_k)_n^2(\eta) = n\eta(1 - \eta) \left[1 + \frac{a}{n} \frac{\varphi_{n-2}(a)}{\varphi_n(a)} \right]. \tag{71}$$

In this case, we check trivially that for $a = 0$ the behavior of σ_k is ordinary binomial. On the other hand, at large n ,

$$\frac{\varphi_{n-2}(a)}{\varphi_n(a)} = \frac{\varphi_{n-2}(a)}{\varphi_{n-1}(a)} \frac{\varphi_{n-1}(a)}{\varphi_n(a)} \sim \frac{n}{a},$$

so the behavior of σ_k is “quasi” ordinary binomial at large n

$$(\sigma_k)_n(\eta) \sim \sqrt{2n\eta(1 - \eta)}. \tag{72}$$

Also in this example the Leibniz triangle rule is obeyed only asymptotically as $n \rightarrow \infty$.

VI. NONLINEAR COHERENT STATES

We introduce here families of coherent states built from the polynomials q_n , following the construction procedure explained in Ref. 5. The basic idea of this section is to view the polynomials $q_n(\eta)$ as *deformations* of η^n , associated to a deformation of the usual factorial $n!$. Although it seems natural to consider the $x_n!$ as a good candidate for the $n!$ -deformation, this choice does not allow in general to obtain the so-called resolution of the identity. To avoid this problem, we need in fact to introduce a new deformation of integers based on integrals.

In the following, we restrict ourselves to the subset $\Sigma_+^{(1)} \subset \Sigma_+$ of generating functions $\Sigma_+^{(1)} = \{e^{\sum_{n=1}^{\infty} a_n t^n} \mid a_1 = 1, a_n \geq 0, \sum_{n=1}^{\infty} a_n < \infty\}$. We do not miss anything by using $\Sigma_+^{(1)}$. Indeed, starting from any function $F(t) = \sum_{n=1}^{\infty} a_n t^n \in \Sigma_0$ with a radius of convergence R , the function $\mathcal{N} = e^{F_0}$ with $F_0(t) = \alpha F(t/\alpha)/a_1$ and $\alpha > R$ belongs to $\Sigma_+^{(1)}$. Now if $\mathcal{N} \in \Sigma_+^{(1)}$, the first element of the sequence $\{x_n\}$ is $x_1 = 1$, and we know that the value $\mathcal{N}(1)$ is well-defined.

Starting from a generating function $\mathcal{N} = e^F \in \Sigma_+^{(1)}$ and the corresponding polynomials q_n , we define

$$f_n = \int_0^\infty q_n(\eta) e^{-\eta} d\eta \quad \text{and} \quad b_{m,n} = \int_0^1 q_m(\eta) q_n(1-\eta) d\eta. \quad (73)$$

The f_n and $b_{m,n}$ are deformations of the usual factorial and beta function deduced from their usual integral definitions by the substitution $\eta^n \mapsto q_n(\eta)$. We prove in Appendix E (for $\mathcal{N} \in \Sigma_+^{(1)}$) the following properties:

$$x_n! \leq f_n, \forall \eta \in \mathbb{R}^+, \sum_{n=0}^\infty \frac{q_n(\eta)}{f_n} < \infty \quad \text{and} \quad b_{m,n} \geq \frac{x_m! x_n!}{(m+n+1)!}. \quad (74)$$

We introduce the function $N(z)$ defined on \mathbb{C} as

$$\forall z \in \mathbb{C}, N(z) = \sum_{n=0}^\infty \frac{q_n(z)}{f_n}. \quad (75)$$

This definition makes sense since from Eq. (74)

$$\sum_{n=0}^\infty \left| \frac{q_n(z)}{f_n} \right| \leq \sum_{n=0}^\infty \frac{q_n(|z|)}{f_n} < \infty. \quad (76)$$

A. Nonlinear coherent states on the complex plane

Let \mathcal{H} be some separable Hilbert space with orthonormal basis $\{|e_n\rangle\}$, $n \in \mathbb{N}$. We define the (normalized) coherent state $\{|z\rangle\}$, $z \in \mathbb{C}$, as

$$|z\rangle = \frac{1}{\sqrt{N(|z|^2)}} \sum_{n=0}^\infty \frac{1}{\sqrt{f_n}} \sqrt{q_n(|z|^2)} e^{i n \arg(z)} |e_n\rangle. \quad (77)$$

These states verify the following resolution of the unity 1_d of \mathcal{H} :

$$\int_{\mathbb{C}} \frac{d^2z}{\pi} e^{-|z|^2} N(|z|^2) |z\rangle \langle z| = 1_d. \quad (78)$$

These coherent states are a natural generalization of the usual harmonic coherent states that correspond to the special polynomials $q_n(\eta) = \eta^n$. The $q_n(\eta) = \eta^n$ are associated to the generating function $\mathcal{N}(t) = e^t \in \Sigma_+^{(1)}$ that gives the usual binomial distribution.

B. Nonlinear spin coherent states

These states can be considered as generalizing the spin coherent states $|\theta, \phi\rangle$ in the $2j + 1$ -dimensional space Hilbert space $\mathcal{H}_j \sim \mathbb{C}^{2j+1}$ with $2j = n$. To each point in the unit sphere S^2 with spherical coordinates (θ, ϕ) , $0 \leq \theta \leq \pi$, $0 \leq \phi < 2\pi$, there corresponds the unit norm state

$$|\theta, \phi\rangle = \frac{1}{\sqrt{\varpi(\theta)}} \sum_{m=-j}^j \sqrt{\frac{q_{j-m}(\cos^2 \theta/2) q_{j+m}(\sin^2 \theta/2)}{b_{j-m, j+m}}} e^{im\phi} |j, m\rangle, \quad (79)$$

where the $b_{m,n}$ are defined in Eq. (73) and $\varpi(\theta)$ is given by

$$\varpi(\theta) = \sum_{m=-j}^j \frac{q_{j-m}(\cos^2 \theta/2) q_{j+m}(\sin^2 \theta/2)}{b_{j-m, j+m}}. \quad (80)$$

Here, $\{|j, m\rangle, -j \leq m \leq j\}$ is the familiar quantum angular momentum orthonormal basis of \mathcal{H}_j (actually it could be any orthonormal basis). Now the family of states (79) resolves the

unity 1_d in \mathcal{H}_j

$$\frac{1}{4\pi} \int_{S^2} \sin \theta \, d\theta \, d\phi \, \varpi(\theta) |\theta, \phi\rangle \langle \theta, \phi| = 1_d. \quad (81)$$

These states can be named coherent states (in a wide sense) which are deformations of the spin coherent states.

VII. CONCLUSION

In previous articles, we have presented a method to construct deformed binomial and Poisson distributions which can be interpreted as probabilities as well. There, the deformed distributions did not preserve the symmetry between *win* and *loss*. In this paper, we studied a normalized *win-loss* symmetrical generalization of the binomial distribution that fulfills the non-negative condition and can be interpreted as the probability associated to a sequence of correlated trials. The related generating functions were defined in a general way and the non-negativeness problem was solved for the simplest class. Three examples were presented; they are interesting at various levels, not the least being that they lead to manageable results. Also, they pertain to a class of what we defined as *logarithmic scale invariant generating functions*. The logarithmic scale invariance is a particularly interesting property as it points to an interpretation of the primary generating function as the number of states of a system where correlations are present. In one of the presented examples, namely, the q -exponential, the Leibniz triangle rule is exactly obeyed and therefore the extensive entropy is Boltzmann-Gibbs. We proved that this is the only case and this result itself is new and rather interesting. In many other cases, the Leibniz rule is only asymptotically true. In a forthcoming paper, we will continue a systematic study of the entropy behavior for various types of deformations of the binomial distribution and their limit distributions.

We also introduced a new class of coherent states built from the symmetric deformed distributions. In future works, we think of using these coherent states to quantize the complex plane of the unit sphere in the spirit of Ref. 5.

Finally, we note that there is another possible generalization of the binomial distribution along the same lines, in which the probabilities of *win* and *loss* are given, respectively, by two different polynomials, according to

$$p_k^{(n)}(\eta) = \frac{x_n!}{x_{n-k}! x_k!} p_k(\eta) q_{n-k}(1-\eta). \quad (82)$$

ACKNOWLEDGMENTS

H. Bergeron and J. P. Gazeau thanks to CBPF and CNPq for financial support and CBPF for hospitality. E.M.F. Curado acknowledges CNPq and FAPERJ for financial support.

APPENDIX A: PROOF OF THE PROPERTIES OF q_n

Let us define $G(t, \eta) = \mathcal{N}(t)^\eta$.

- (a) We remark first that $G(0, \eta) = 1$ (because $\mathcal{N}(0) = 1$), then $q_0(\eta) = 1$. Moreover, $\frac{\partial G}{\partial t}(0, \eta) = \eta \mathcal{N}'(0) G(0, \eta - 1) = \eta/x_1$, therefore $q_1(\eta) = \eta$.
More generally we have $\frac{\partial G}{\partial t}(t, \eta) = \eta \mathcal{N}'(t) G(t, \eta - 1)$. Identifying the series expansion in t of the left and right hand sides, we obtain the sought recurrence relation (21).
- (b) Knowing $q_0(\eta) = 1$ and the recurrence relation (21), we deduce that the functions q_n are polynomials, the degree of q_n being n . Furthermore, the recurrence relation shows that $\forall n \neq 0$, $q_n(0) = 0$ and the series expansion in t of $G(t, 1)$ shows that $\forall n \in \mathbb{N}$, $q_n(1) = 1$. Finally, we remark that $G(t, \eta)G(t, 1 - \eta) = \mathcal{N}(t)$. Using the series expansion in t of the left hand side of this equation, we obtain the normalization condition (9).

- (c) More generally since the series expansion in t of $G(t, \eta)$ is valid for any $\eta \in \mathbb{C}$ and since $G(t, z_1)G(t, z_2) = G(t, z_1 + z_2)$, we obtain the functional relation (23).

APPENDIX B: COMBINATORIAL EXPRESSION OF q_n

In the case where the generating function of polynomials q_n is $(\mathcal{N}(t))^\eta$ the application of Faà di Bruno’s formula (0.430-2) in Ref. (17) to (23) leads to the complete (although not so easily manageable) expression for polynomials q_n , for $n \geq 1$

$$q_n(\eta) = x_n! \sum_{m=1}^n \frac{\Gamma(\eta + 1)}{\Gamma(\eta - m + 1)} \sum_{\mathcal{I}(i_s)} \frac{1}{i_1! i_2! \dots i_k!} \frac{1}{\prod_{s=1}^k (x_s!)^{i_s}}, \tag{B1}$$

where the multiple summation symbol $\mathcal{I}(i_s)$ means summation on all possible variable-length -uples of nonnegative integers i_1, i_2, \dots, i_k , submitted to the 2 constraints

$$\sum_{s=1}^k s i_s = n,$$

$$\sum_{s=1}^k i_s = m.$$

APPENDIX C: PROOF OF THE NON-NEGATIVENESS OF q_n

Definition C.1. Using the notation $G_{\mathcal{N},\eta}(t) = \mathcal{N}(t)^\eta$, Σ_+ is defined as

$$\tilde{\Sigma}_+ = \{ \mathcal{N} \in \Sigma \mid \forall \eta \in]0, 1], G_{\mathcal{N},\eta} \in \Sigma \}. \tag{C1}$$

Proposition C.1. The set $\tilde{\Sigma}_+$ defined in (C1) satisfies the following properties:

1. $\forall \mathcal{N}_1, \mathcal{N}_2 \in \Sigma_+, \mathcal{N}_1 \mathcal{N}_2 \in \tilde{\Sigma}_+$
2. $\forall F \in \Sigma_0, t \mapsto e^{F(t)} \in \tilde{\Sigma}_+$ (the set Σ_0 has been introduced in the sequence just before (7)).

Proof. To prove point (1), from Proposition 2.2 of Ref. 2 we have that \mathcal{N}_1 and $\mathcal{N}_2 \in \Sigma$ implies that $\mathcal{N}_1 \mathcal{N}_2 \in \Sigma$. Furthermore, $G_{\mathcal{N}_1 \mathcal{N}_2, \eta} = G_{\mathcal{N}_1, \eta} G_{\mathcal{N}_2, \eta}$, then if \mathcal{N}_1 and $\mathcal{N}_2 \in \tilde{\Sigma}_+$, by definition $G_{\mathcal{N}_1, \eta}$ and $G_{\mathcal{N}_2, \eta} \in \Sigma$. Then from Proposition 2.2 of Ref. 2, we deduce $G_{\mathcal{N}_1 \mathcal{N}_2, \eta} \in \Sigma$.

The part (2) is obtained always from Proposition 2.2 of Ref. 2. First, if $F \in \Sigma_0$, then $\mathcal{N} = e^F \in \Sigma$ (Proposition 2.2). Furthermore, $G_{\mathcal{N}, \eta}(t) = e^{\eta F(t)}$. For $\eta \in]0, 1]$, $t \mapsto \eta F(t)$ also belongs to Σ_0 then we deduce always from Proposition 2.2, $G_{\mathcal{N}, \eta} \in \Sigma$. □

1. The characterization of $\tilde{\Sigma}_+$

In fact, point (2) in Proposition C.1 is not only a sufficient but also a necessary condition to obtain functions of $\tilde{\Sigma}_+$. In order to prove it, we start from the following lemma:

Lemma C.1. For any $\mathcal{N} \in \tilde{\Sigma}_+$, the polynomials q_n verify $q'_n(0) \geq 0$ and furthermore

$$\frac{\mathcal{N}'(t)}{\mathcal{N}(t)} = \sum_{n=0}^{\infty} \frac{(n+1)q'_{n+1}(0)}{x_{n+1}!} t^n. \tag{C2}$$

Proof. To begin with, $q_0(\eta) = 1$, so $q'_0(\eta) = 0$. For $n \geq 1$, the polynomials $q_n(\eta)$ are nonnegative on the interval $\eta \in (0, 1]$ and $q_n(0) = 0$. Then $q'_n(0)$ cannot be negative, otherwise $q_n(\eta)$ would be negative on some interval $(0, \epsilon)$. We conclude that $\forall n \in \mathbb{N}, q'_n(0) \geq 0$.

Now using the function $G_{\mathcal{N},\eta}(t) = \mathcal{N}(t)^\eta$ and by a differentiation with respect to t , we obtain first

$$\forall \eta \in (0, 1], \mathcal{N}'(t)\mathcal{N}(t)^{(\eta-1)} = \frac{1}{\eta} \sum_{n=0}^{\infty} \frac{(n+1)q_{n+1}(\eta)}{x_{n+1}!} t^n. \tag{C3}$$

Taking into account the relation $\forall n \neq 0, q_n(0) = 0$, and assuming that the permutation between $\lim_{\eta \rightarrow 0}$ and \sum is valid we obtain Eq. (C2). \square

This lemma leads to the theorem

Theorem C.1. *For all $\mathcal{N} \in \tilde{\Sigma}_+$, $\ln \mathcal{N} \in \Sigma_0$ and*

$$\ln \mathcal{N}(t) = \sum_{n=1}^{\infty} \frac{q'_n(0)}{x_n!} t^n. \tag{C4}$$

Furthermore, we deduce the characterization

$$\tilde{\Sigma}_+ = \{e^F \mid F \in \Sigma_0\}. \tag{C5}$$

Proof. Equation (C4) is immediately obtained by integrating Eq. (C2) term by term and taking into account the value $\ln \mathcal{N}(0) = 0$. Using the property $q'_n(0) \geq 0$, which was shown in Lemma C.1, and the fact that $q'_1(0) = 1$, as $q_1(\eta) = \eta$, we deduce that $\ln \mathcal{N} \in \Sigma_0$.

The last part of the theorem results from Proposition C.1 (point (2)) and the previous comment. \square

Theorem C.1 solves completely the problem of generating functions leading to symmetric deformations of the binomial distribution preserving a statistical interpretation.

APPENDIX D: EXACT SOLUTION TO THE LEIBNIZ TRIANGLE RULE

Here, we prove a uniqueness result about the Leibniz triangle rule

$$\varpi_k^{n-1} = \varpi_k^n + \varpi_{k+1}^n, \quad 0 \leq k \leq n-1. \tag{D1}$$

Proposition D.1. *With the notations of Sec. III, given a sequence of numbers (x_n) with $x_0 = 0$, $x_n > 0$ for all $n > 0$, with generating function $\mathcal{N}(t) = \sum_{n=0}^{\infty} t^n/x_n!$, and the polynomials $q_n(\eta)$ with generating function $(\mathcal{N}(t))^\eta = \sum_{n=0}^{\infty} q_n(\eta)t^n/x_n!$, the quantities*

$$\frac{\mathfrak{P}_k^{(n)}(\eta)}{\binom{n}{k}} = \binom{y_n}{y_k} q_k(\eta)q_{n-k}(1-\eta), \tag{D2}$$

where $y_n = x_n/n$, are solutions to (D1) for $n \geq 1$ if and only if the polynomials q_n and the sequence (y_n) read as either

$$q_n(\eta) = \prod_{k=1}^n \frac{k-1+\alpha\eta}{k-1+\alpha}, \quad y_n = \frac{\alpha}{n+\alpha-1}, \quad n \geq 1, \tag{D3}$$

where α is an arbitrary constant (finite case), or

$$q_n(\eta) = \eta^n, \quad y_n = 1, \quad n \geq 1, \tag{D4}$$

(infinite case).

Proof. After injection of the expression (D2) into the two members of the equality (D1) and simplification of the factorials from both sides, we obtain for all $n \geq 1$ and $0 \leq k \leq n-1$

$$q_k(\eta)q_{n-k-1}(1-\eta) = \frac{y_n}{y_{n-k}} q_k(\eta)q_{n-k}(1-\eta) + \frac{y_n}{y_{k+1}} q_{k+1}(\eta)q_{n-k-1}(1-\eta). \tag{D5}$$

We now particularize this identity to the case $k = 0$. From Sec. III, we know that $q_0(\eta) = 1$ and $q_1(\eta) = \eta$. There follows the recurrence relation

$$q_n(\eta) = \left(1 + \frac{y_n}{y_1}(\eta - 1)\right) q_{n-1}(1 - \eta) \quad (\text{D6})$$

and the explicit solution

$$q_n(\eta) = \prod_{k=1}^n \left(1 + \frac{y_k}{y_1}(\eta - 1)\right). \quad (\text{D7})$$

We now divide (D5) by $q_k(\eta)$ and put $\eta = 0$. From Sec. III, we know that $q_n(1) = 1$ for all n and from (D6) we have $q_{k+1}(0)/q_k(0) = 1 - y_{k+1}/y_1$. Hence, we obtain the functional equation for $Z(n) := 1/y_n$

$$Z(n) = Z(n - k) + Z(k + 1) - Z(1). \quad (\text{D8})$$

This equation is of Cauchy-Pexider type²⁰ and is easily solved by recurrence. Its general solution is $Z(n) = \mu n + \nu$, where μ and ν are arbitrary constants. Now, we note that the only ratios y_k/y_1 appear in the solution (D7). Therefore, for all $k \geq 1$,

$$\frac{y_k}{y_1} = \frac{\mu + \nu}{\mu k + \nu} = \frac{1}{\frac{\mu}{\mu + \nu}k + \frac{\nu}{\mu + \nu}} = \frac{\alpha}{k - 1 + \alpha}, \quad (\text{D9})$$

where we have introduced the parameter $\alpha = 1 + \nu/\mu$. With this expression at hand, we eventually obtain (D3). The limit case (D4) corresponds to $\alpha \rightarrow \infty$ and is trivially verified. \square

APPENDIX E: PROOF OF THE PROPERTIES OF $\Sigma_+^{(1)}$

Using Eq. (29), we obtain

$$f_n = \sum_{k=1}^n \frac{B_{n,k}(a_1, a_2, \dots, a_{n-k+1})}{B_n(a_1, a_2, \dots, a_n)} k! \geq \frac{n! B_{n,n}(a_1)}{B_n(a_1, a_2, \dots, a_n)}. \quad (\text{E1})$$

Because $a_1 = 1$ for functions of $\Sigma_+^{(1)}$, and using Eq. (27), we find the first property $f_n \geq x_n!$. Therefore, we have

$$\forall \eta > 0, 0 \leq \sum_{n=0}^{\infty} \frac{q_n(\eta)}{f_n} \leq \sum_{n=0}^{\infty} \frac{q_n(\eta)}{x_n!}, \quad (\text{E2})$$

and the right hand side of the inequality is finite because the radius of convergence of $\mathcal{N}(t)^\eta = \sum_{n=0}^{\infty} \frac{q_n(\eta)}{x_n!} t^n$ is greater than 1 for functions of $\Sigma_+^{(1)}$. We conclude that $\sum_{n=0}^{\infty} \frac{q_n(\eta)}{f_n}$ is finite. Finally, we have

$$b_{m,n} = \int_0^1 q_m(\eta) q_n(1 - \eta) d\eta = \sum_{i=0}^m \sum_{j=0}^n \frac{q_m^{(i)}(0) q_n^{(j)}(0)}{(i + j + 1)!} \geq \frac{q_m^{(m)}(0) q_n^{(n)}(0)}{(m + n + 1)!}. \quad (\text{E3})$$

Since for functions belonging to $\Sigma_+^{(1)}$ we have $q_n^{(n)}(0) = x_n!$, we obtain the last inequality $b_{m,n} \geq x_m! x_n! / (m + n + 1)!$.

¹ E. M. F. Curado, J. P. Gazeau, and L. M. C. S. Rodrigues, "On a generalization of the binomial distribution and its Poisson-like limit," *J. Stat. Phys.* **146**, 264–280 (2012).

² H. Bergeron, E. M. F. Curado, J. P. Gazeau, and L. M. C. S. Rodrigues, "Generating functions for generalized binomial distributions," *J. Math. Phys.* **53**, 103304 (2012).

³ E. M. F. Curado, J. P. Gazeau, and L. M. C. S. Rodrigues, "Non-linear coherent states for optimizing quantum information, Proceedings of the Workshop on Quantum Nonstationary Systems, October 2009, Brasilia. Comment section (CAMOP)," *Phys. Scr.* **82**, 038108 (2010).

⁴ V. V. Dodonov, *J. Opt. B: Quantum Semiclassical Opt.* **4**, R1 (2002), and references therein.

⁵ J.-P. Gazeau, *Coherent States in Quantum Physics* (Wiley, 2009).

⁶ R. Hanel, S. Thurner, and C. Tsallis, *Eur. Phys. J. B* **72**, 263–268 (2009).

- ⁷C. Tsallis, “Nonadditive entropy and nonextensive statistical mechanics: An overview after 20 years,” *Braz. J. Phys.* **39**, 337–356 (2009).
- ⁸B. de Finetti, *Atti Accad. Naz. Lincei, Mem., Cl. Sci. Fis., Mat. Nat.* **4**, 251299 (1931); D. Heath and W. Sudderth, *Am. Stat.* **30**, 188 (1976).
- ⁹J. S. Kim, “Tsallis entropy and entanglement constraints in multiqubit systems,” *Phys. Rev. A* **81**, 062328 (2010).
- ¹⁰In our articles, the terms *generalized binomial* or *binomial-like* or *deformed binomial* distribution refer to mathematical expressions like (2) in which x_n are non-negative real arbitrary numbers. In most papers devoted to probability and statistics in which such terms are used, their meaning is a lot more restrictive and concerns counting combinatorics and integer numbers.
- ¹¹A. D. Polyanin and A. V. Manzhirov, *Handbook of Integral Equations: Exact Solutions (Supplement. Some Functional Equations)* (Faktorial, Moscow, 1998) (in Russian), see also <http://eqworld.ipmnet.ru/en/solutions/fe/fe2104.pdf>.
- ¹²S. Roman, *The Umbral Calculus* (Dover Publications, 2005); R. Mullin and G. C. Rota A, “Theory of binomial enumeration,” in *Graph Theory and its Applications* (Academic Press, New York, 1970); J. Aczél, “Functions of binomial type mapping groupoids into rings,” *Math. Z.* **154**, 115–124 (1977); A. Di Bucchianico, *Probabilistic and Analytical Aspects of the Umbral Calculus*, CWI Tracts Vol. 119 (Centrum voor Wiskunde en Informatica, Amsterdam, 1997), pp. 1–148.
- ¹³L. Comtet, *Advanced Combinatorics* (Reidel, 1974).
- ¹⁴We are grateful to the referee for mentioning this point.
- ¹⁵N. Johnson, S. Kotz, and A. W. Kempf, *Univariate Discrete Distributions*, 2nd ed. (John Wiley and Sons, 1992), p. 239.
- ¹⁶“Pólya distribution,” *Encyclopedia of Mathematics* [URL: http://www.encyclopediaofmath.org/index.php?title=P%C3%B3lya_distribution&oldid=25927].
- ¹⁷I. S. Gradshteyn and I. M. Ryzhik, *Table of Integrals, Series, and Products*, 5th ed. (Academic Press, USA, 1994).
- ¹⁸M. Abramowitz and I. A. Stegun, *Handbook of Mathematical Functions*, National Bureau of Standards, 10th ed., Applied Mathematical Series Vol. 55 (National Bureau of Standards, Washington, 1972).
- ¹⁹H. Bergeron, E. M. F. Curado, J. P. Gazeau, and L. M. C. S. Rodrigues, “A note about combinatorial sequences and incomplete gamma function,” e-print [arXiv:1309.6910v1](https://arxiv.org/abs/1309.6910v1).
- ²⁰J. Aczél and J. Dhombres, *Functional equations in several variables*, Cambridge University Press, New York, 1989.